Causal and statistical inference in the presence of network dependence

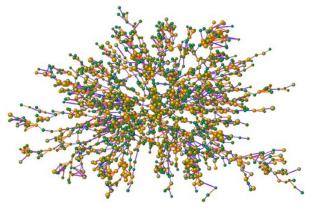
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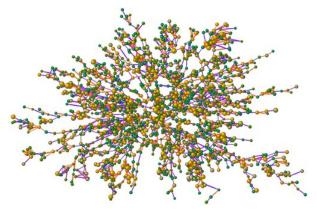
outline

- Brief history of causal inference using network data.
- What is network dependence, and why is it a problem?
- Two partial solutions:
 - 1. Subsample conditionally independent observations
 - naive but easy to understand, implement, and generalize,
 - dependence due to contagion;
 - 2. Semiparametric approach based on the efficient influence function
 - more sophisticated and powerful but less intuitive and difficult to implement.



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- Each node is associated with an outcome, treatment, covariates.
- Causal effects of interest include peer effects, treatment effects, spillover/interference effects, effects of network interventions, ...



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Two challenges for causal inference using network data:

- nonparametric identification of causal effects (interference, confounding by homophily, positivity violations),
- statistical inference in the presence of network dependence.

brief history of causal inference using network data

- Christakis and Fowler (2007, 2008, 2009, 2010, 2011, 2012) initiated a wave of interest in estimating peer effects from social network data.
 - ► To examine peer effects, they fit models

$$Y_{ego}^t \sim Y_{alter}^{t-1}, Y_{alter}^{t-2}, Y_{ego}^{t-2}, \boldsymbol{C}_{ego}$$

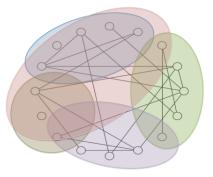
- Widely publicized results include significant peer effects for obesity, smoking, alcohol consumption, sleep habits, etc.
- Researchers began using similar models to assess peer effects across a wide range of disciplines and problems (e.g. Ali and Dwyer, 2009; Cacioppo et al., 2009; 2008; Lazer et al., 2010; Rosenquist et al., 2010, Wasserman 2012).

brief history causal inference using network data

- There is growing interest in randomization-based inference for networks (e.g. Toulis & Kao, 2013; Bowers et al., 2013; Aronow & Samii, 2013; Eckles et al., 2014, Choi 2016).
- Work on interference usually relies on randomization and on the assumption of partial interference, but may provide a solution to the problem of network dependence in cluster randomized trials (e.g. Sobel, 2006; Hong & Raudenbush, 2006; Rosenbaum, 2007; Hudgens & Halloran, 2008; Tchetgen Tchetgen & VanderWeele, 2012; Liu & Hudgens, 2014).
- Mathematical modeling of contagious processes avoids these problems but is highly dependent on parametric assumptions about agent-based processes (e.g. Steglich, Snijders & Pearson, 2007; Railsback & Grimm, 2011).

sources of network dependence

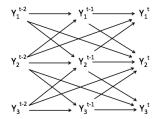
 Latent variables cause outcomes among close social contacts to be more correlated than among distant contacts. (E.g. homophily, geography, shared culture, shared genetics.)



Similar to spatial dependence.

sources of network dependence

▶ **Contagion** implies information barrier structures, e.g. $[Y_1^t \perp Y_2^t \mid Y_1^{t-2}, Y_2^{t-2}, Y_1^{t-1}, \text{ and } Y_2^{t-1}]$ and $[Y_1^{t-2} \perp Y_3^{t-1}]$.



- When a network is observed at a single time point, this will resemble latent variable dependence.
- If the network is observed frequently, so that the outcome can't diffuse very far between observations, we can harness conditional independence restrictions to facilitate inference.

why is dependence a problem?

- Statistical analysis that incorrectly assumes independence will be invalid.
- This is a very hard problem when dependence is due to latent variables and is unstructured.
- It's not quite as hard when dependence is due to contagion.
- Two problems for traditional frequentist inference:
 - CLT may not hold,
 - Standard error estimates and resulting inference will be anticonservative.

 $\blacktriangleright\,$ If $\bar{Y} \rightarrow \mu,$ the rate of convergence is determined by

$$var(\bar{Y}) = \frac{1}{n^2} \left\{ \sum_{i=1}^n \sigma^2 + \sum_{i \neq j} cov(Y_i, Y_j) \right\}$$

$$b_n = \frac{1}{n} \sum_{i \neq j} cov(Y_i, Y_j)$$

$$var(\bar{Y}) = rac{\sigma^2}{n/\left(1 + rac{b_n}{\sigma^2}
ight)}$$

▶ If a CLT holds, then

$$\sqrt{\frac{n}{1+\frac{b_n}{\sigma^2}}}\left\{\bar{Y}-\mu\right\}\stackrel{d}{\rightarrow}\mathsf{N}(0,\sigma^2)$$

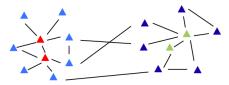
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naive solution

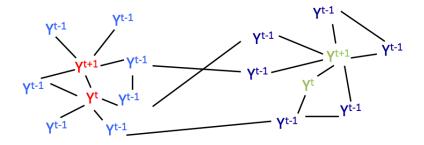
joint work with Tyler VanderWeele

- Summary: Create conditionally independent units; analyze with standard, i.i.d. models, but conditional on "information barriers."
- Randomly sample non-overlapping groups from the network.



This will allow us to condition on an "information barrier."

- Now we can estimate conditional estimands using standard statistical machinery like GLMs.
 - The residuals will be uncorrelated across subjects despite the dependence structure.



• Regress Y^{t+1} on Y^t conditional on $\{Y^{t-1}\}$

For details see Ogburn & VanderWeele, Vaccines, contagion, and social networks (forthcoming in AoAS)

naive solution

Pros

- easy to understand, easy to implement
- generalizable to many estimands and models (in principle)
- may be feasible if full data structure is unavailable, as long as information barriers can be found
- Cons
 - dependence due to contagion only
 - sample size < true effective sample size</p>
 - requires throwing away data
 - Iow power
 - estimand must be conditional
 - more appropriate for causal effects than for sample means

more principled solution

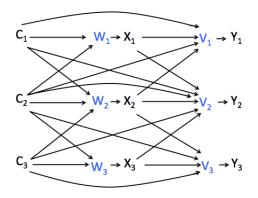
joint work with Oleg Sofrygin, Ivan Diaz, Mark van der Laan

- Extension of semiparametric, influence-function-based inference from the iid setting.
- We define a model *M*, which restricts the observed data distribution in some way(s).
- We are interested in estimating a parameter ψ under model *M*, i.e. a functional of the observed data.
- Under \mathcal{M} , there is a class of **influence functions** for ψ .
 - Each (RAL) estimator $\hat{\psi}$ is paired with an IF ϕ , and in the iid setting

$$\sqrt{n}(\hat{\psi}-\psi) \stackrel{p}{\approx} \frac{1}{\sqrt{n}} \sum_{i=1}^{n} \varphi(O_i)$$

• Because the IF has mean 0 at the true parameter value, we can use it to create unbiased estimating functions for ψ .

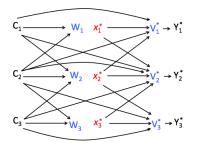
- van der Laan (2014) extended this approach to settings with interference and/or contagion.
 - Not partial interference, but each subject can only interfere with $\leq K$ other subjects.
- ▶ We extend van der Laan (2014) to social network settings:
 - K grows with n
 - highly connected "hubs" may exert undo influence
 - estimation of causal effects of interventions on features of network topology
- This framework can handle longitudinal data, but for simplicity we focus on the single-time-point setting.



We make independence assumptions that entail

- there is no unmeasured confounding,
- $C_i \perp C_j$ if *i* and *j* have no friends in common,
- ► $Y_i \perp Y_j | \text{parents}$ and $X_i \perp X_j | \text{parents}$ if *i* and *j* have no friends in common.

The simplest kind of intervention deterministically sets X to a user-specified value x*:



- Y^{*}_i is the counterfactual outcome of individual i in a hypothetical world in which P(X = x^{*}) = 1.
 - Peer effects: X_i could be a function of alters' outcomes at a previous time point.
- We are interested in $E\left[\bar{Y}^*\right]$, where $\bar{Y}^* = \frac{1}{n} \sum_{i=1}^{n} Y_i^*$.

• $E\left[\bar{Y}^*\right]$ is identified by the parameter

$$\psi = \frac{1}{n} \sum_{i=1}^{n} E\left[\sum_{y} y \, p_{Y}(y | V_{i}^{*})\right] = \frac{1}{n} \sum_{i=1}^{n} \sum_{v} \left[\sum_{y} y \, p_{Y}(y | v)\right] P\left[V_{i}^{*} = v\right].$$

 The efficient influence function for ψ (in a particular semiparametric model) is

$$\varphi(\mathbf{O}) = \sum_{j=1}^{n} \frac{1}{n} \sum_{i=1}^{n} E\left[\sum_{y} y \, p_{Y}(y | V_{i}^{*}) \mid C_{j} = c_{j}\right] - \psi \\ + \frac{1}{n} \sum_{i=1}^{n} \frac{\frac{1}{n} \sum_{j=1}^{n} P(V_{j}^{*} = v_{i})}{\frac{1}{n} \sum_{j=1}^{n} P(V_{j} = v_{i})} \left\{ y_{i} - \sum_{y} y \, p_{Y}(y | v_{i}) \right\}$$

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• This is of the form $\frac{1}{n}\sum_{i=1}^{n}\varphi_i(\mathbf{O})$ instead of $\frac{1}{n}\sum_{i=1}^{n}\varphi(O_i)$.

- Turning the efficient IF into an estimating equation and solving it gives us an estimate ψ of ψ.
- $ilde{\psi}$ is asymptotically efficient and doubly robust.
- ► If each subject interferes with ≤ K other subjects, as in van der Laan (2014), then

$$\sqrt{n}(\tilde{\psi} - \psi)
ightarrow N(0, var(IF))$$

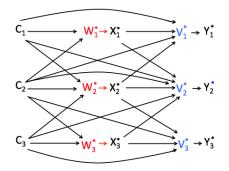
▶ Instead, we let $K_n \to \infty$ as $n \to \infty$ s.t. $\frac{K_n^2}{n} \to 0$. Then

$$\sqrt{C_n}(\hat{\psi}-\psi) \rightarrow N(0, var(IF)),$$

where $\frac{n}{K_n^2} \leq C_n \leq n$.

stochastic network interventions

We can also identify the effects of stochastic interventions that replace f_X with a new, user-specified distribution:



For each x in the support of X, X_i is set by the intervention to x with probability given by the stochastic intervention distribution.

stochastic network interventions

- Examples include
 - interventions that add, remove, or relocate ties in the network.
 - interventions that change the dependence of a subject's treatment on other subjects' covariates, or of a subject's outcome on other subjects' covariates and treatments.
 - Interventions on summary features of network topology:
 - An intervention on features of the network topology replaces T with the members of a class 𝔅^{*} of n × n adjacency matrices that share the intervention features, stochastically according to some probability distribution 𝔅^{T*} over 𝔅^{*}.
 - Whether or not we can define, identify, and estimate interventions involving these features of network topology hinges crucially on the positivity assumption.
 - e.g. degree / centrality

principled approach

Pros

- uses all of the available data
- estimands are unconditional
- efficient and doubly robust estimation
- Cons
 - hard(er) to understand, hard to implement
 - ▶ may not be clear in finite samples what to do with K and with hubs

For details see Ogburn et al, Causal inference for social network data (available on arXiv)

summary and next steps

- Although it is accepted practice in many areas, it can be very dangerous to assume that observations are independent when they may not be!
- When it's available, we can use the information barrier structure to facilitate inference even when subjects are connected in complicated ways.
- Future work is needed to adapt results from spatial statistics to deal with non-independence of observations.
 - This is necessary for latent variable dependence.
 - It is desirable when network dependence is due to contagion, because it permits inference from more realistic/feasible data structures.

Thank you



Why can't we use spatial dependence results?

- Network topology doesn't naturally correspond to Euclidean space.
 - In order to embed a network in \mathbb{R}^d , we would have to let d grow with sample size.
 - Spatial results require d to be fixed.
- Population growth is usually assumed to occur at the boundaries of the *d*-dimensional space.
 - It's not clear how to define boundaries in networks.
- Mixing assumptions and m-dependence don't imply bounded correlation structure.
 - In spatial data most observations are distant from one another.
 - The maximum network-based distance between two observations may be very small.
 - The distance distribution may not be right-skewed enough.